

Supplementary information

A scaling law of multilevel evolution: how the balance between within- and among-collective evolution is determined

Nobuto Takeuchi^{*1,4}, Namiko Mitarai^{2,4}, and Kunihiro Kaneko^{3,4}

¹School of Biological Sciences, University of Auckland, Private Bag 92019, Auckland 1142, New Zealand

²The Niels Bohr Institute, University of Copenhagen, Blegdamsvej 17, Copenhagen, 2100-DK, Denmark

³Graduate School of Arts and Sciences, University of Tokyo, Komaba 3-8-1, Meguro-ku, Tokyo 153-8902, Japan

⁴Research Center for Complex Systems Biology, Universal Biology Institute, University of Tokyo, Komaba 3-8-1, Meguro-ku, Tokyo 153-8902, Japan

1 Derivation of Eq. (3)

In this section, we derive Eq. (3) of the main text, which is redisplayed below:

$$\mathbb{E}[\Delta\langle\langle k_{ij} \rangle\rangle] = \langle\langle w_{ij} \rangle\rangle^{-1} \left\{ \text{cov}_i [\langle w_{ij} \rangle, \langle k_{ij} \rangle] + \text{ave}_i [\text{cov}_{ij} [w_{ij}, k_{ij}]] \right\},$$

where the symbols are defined as follows:

$$\langle\langle k_{ij} \rangle\rangle := \frac{1}{M} \sum_{i=1}^L n_i \langle k_{ij} \rangle, \tag{S1}$$

^{*}Corresponding author. E-mail: nobuto.takeuchi@auckland.ac.nz

where M is the total number of replicators, n_i is the number of replicators in collective i , L is the number of collectives, and

$$\begin{aligned}
\langle k_{i\tilde{j}} \rangle &:= \frac{1}{n_i} \sum_{j=1}^{n_i} k_{ij}, \\
\langle\langle w_{i\tilde{j}} \rangle\rangle &:= \frac{1}{M} \sum_{i=1}^L n_i \langle w_{i\tilde{j}} \rangle, \\
\langle w_{i\tilde{j}} \rangle &:= \frac{1}{n_i} \sum_{j=1}^{n_i} w_{ij}, \\
\text{cov}_{i\tilde{j}} [\langle w_{i\tilde{j}} \rangle, \langle k_{i\tilde{j}} \rangle] &:= \frac{1}{M} \sum_{i=1}^L n_i (\langle w_{i\tilde{j}} \rangle - \langle\langle w_{i\tilde{j}} \rangle\rangle) (\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle), \\
\text{cov}_{i\tilde{j}} [w_{ij}, k_{ij}] &:= \frac{1}{n_i} \sum_{j=1}^{n_i} (w_{ij} - \langle w_{i\tilde{j}} \rangle) (k_{ij} - \langle k_{i\tilde{j}} \rangle), \\
\text{ave}_{i\tilde{j}} [\text{cov}_{i\tilde{j}} [w_{ij}, k_{ij}]] &:= \frac{1}{M} \sum_{i=1}^L n_i \text{cov}_{i\tilde{j}} [w_{ij}, k_{ij}].
\end{aligned} \tag{S2}$$

In each generation, a replicator is sampled M times with replacement from replicators of the previous generation with probabilities proportional to fitness w_{ij} , as in the Wright-Fisher process (see the main text under “Model”). To express $\langle\langle k_{i\tilde{j}} \rangle\rangle$ in the next generation, we introduce the following symbols. Let I_l be the index of the collective to which the l th sampled replicator belongs ($l \in \{1, 2, \dots, M\}$), J_l be the index of the sampled replicator within collective I_l , and $P(I_l = i, J_l = j)$ be the probability that replicator j in collective i is sampled. By the definition of the Wright-Fisher process,

$$P(I_l = i, J_l = j) = \frac{w_{ij}}{M \langle\langle w_{i\tilde{j}} \rangle\rangle}. \tag{S3}$$

Moreover, let $\epsilon_{I_l J_l}$ be the effect of mutation and $P(\epsilon_{I_l J_l})$ be its probability distribution function ($\epsilon_{I_l J_l}$ takes a value of 0 with a probability $1 - m$ or a value sampled from a Gaussian distribution with mean 0 and variance σ with a probability m). Finally, let $\mathbb{E}[x]$ denote the expected value of x after one iteration of the Wright-Fisher process; e.g.,

$$\mathbb{E}[k_{I_l J_l} + \epsilon_{I_l J_l}] = \sum_{i=1}^L \sum_{j=1}^{n_i} P(I_l = i, J_l = j) \int dP(\epsilon_{I_l J_l}) (k_{ij} + \epsilon_{I_l J_l}). \tag{S4}$$

Using these definitions, we can express the expected change of $\langle\langle k_{i\tilde{j}} \rangle\rangle$ per generation, denoted by $\mathbb{E}[\Delta \langle\langle k_{i\tilde{j}} \rangle\rangle]$, as follows:

$$\mathbb{E}[\Delta \langle\langle k_{i\tilde{j}} \rangle\rangle] = \mathbb{E} \left[M^{-1} \sum_{l=1}^M (k_{I_l J_l} + \epsilon_{I_l J_l}) \right] - \langle\langle k_{i\tilde{j}} \rangle\rangle. \tag{S5}$$

Since I_l and J_l are independent and identically distributed for different values of l , we can remove the summation in the above equation to obtain

$$\begin{aligned}
\mathbb{E}[\Delta \langle\langle k_{i\tilde{j}} \rangle\rangle] &= \mathbb{E}[k_{IJ} + \epsilon_{IJ}] - \langle\langle k_{i\tilde{j}} \rangle\rangle \\
&= \mathbb{E}[k_{IJ}] - \langle\langle k_{i\tilde{j}} \rangle\rangle,
\end{aligned} \tag{S6}$$

where we used the fact that $\mathbb{E}[\epsilon_{IJ}] = 0$ and omitted subscript l .

The first term on the RHS of Eq. (S6) can be calculated as follows:

$$\begin{aligned}
\mathbb{E}[k_{IJ}] &= \sum_{i=1}^L \sum_{j=1}^{n_i} P(I=i, J=j) k_{ij} \\
&= \sum_{i=1}^L \sum_{j=1}^{n_i} \frac{w_{ij}}{M \langle\langle w_{i\tilde{j}} \rangle\rangle} k_{ij} \\
&= \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \frac{1}{M} \sum_{i=1}^L \sum_{j=1}^{n_i} w_{ij} k_{ij} \\
&= \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \frac{1}{M} \sum_{i=1}^L n_i \frac{1}{n_i} \sum_{j=1}^{n_i} w_{ij} k_{ij} \\
&= \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \frac{1}{M} \sum_{i=1}^L n_i (\langle w_{i\tilde{j}} \rangle \langle k_{i\tilde{j}} \rangle + \text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}]) \\
&= \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \left\{ \frac{1}{M} \left(\sum_{i=1}^L n_i \langle w_{i\tilde{j}} \rangle \langle k_{i\tilde{j}} \rangle \right) + \text{ave}_{\tilde{i}} [\text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}]] \right\} \\
&= \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} (\text{cov}_{\tilde{i}}[\langle w_{i\tilde{j}} \rangle, \langle k_{i\tilde{j}} \rangle] + \langle\langle w_{i\tilde{j}} \rangle\rangle \langle\langle k_{i\tilde{j}} \rangle\rangle + \text{ave}_{\tilde{i}} [\text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}]]) \\
&= \langle\langle k_{i\tilde{j}} \rangle\rangle + \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \{ \text{cov}_{\tilde{i}}[\langle w_{i\tilde{j}} \rangle, \langle k_{i\tilde{j}} \rangle] + \text{ave}_{\tilde{i}} [\text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}]] \}.
\end{aligned} \tag{S7}$$

Substituting Eq. (S7) into Eq. (S6), we obtain Eq. (3).

2 Derivation of Eq. (4)

In this section, we derive Eq. (4) of the main text, which is redisplayed below:

$$\mathbb{E}[\Delta \langle\langle k_{i\tilde{j}} \rangle\rangle] = s_a v_a - s_w v_w + O(s_w^2) + O(s_a^2),$$

where the symbols are defined as follows:

$$\begin{aligned}
v_a &:= \text{cov}_{\tilde{i}}[\langle k_{i\tilde{j}} \rangle, \langle k_{i\tilde{j}} \rangle] = \frac{1}{M} \sum_{i=1}^L n_i (\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle)^2 \\
v_{wi} &:= \text{cov}_{i\tilde{j}}[k_{ij}, k_{ij}] = \frac{1}{n_i} \sum_{j=1}^{n_i} (k_{ij} - \langle k_{i\tilde{j}} \rangle)^2, \\
v_w &:= \text{ave}_{\tilde{i}}[v_{wi}] = \frac{1}{M} \sum_{i=1}^L n_i v_{wi}.
\end{aligned} \tag{S8}$$

Equation (4) is obtained by expanding $\langle w_{i\tilde{j}} \rangle$ and w_{ij} in Eq. (3), i.e.,

$$\Delta \langle\langle k_{i\tilde{j}} \rangle\rangle = \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \text{cov}_{\tilde{i}}[\langle w_{i\tilde{j}} \rangle, \langle k_{i\tilde{j}} \rangle] + \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \text{ave}_{\tilde{i}} [\text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}]],$$

as Taylor series around $\langle k_{i\tilde{j}} \rangle = \langle\langle k_{i\tilde{j}} \rangle\rangle$ and $k_{ij} = \langle k_{i\tilde{j}} \rangle$, respectively.

First, we obtain the first term of Eq. (4), which stems from the first term of Eq. (3). We assume that $\langle w_{i\tilde{j}} \rangle$ is an analytic function of $\langle k_{i\tilde{j}} \rangle$ and that $\langle w_{i\tilde{j}} \rangle = \langle\langle w_{i\tilde{j}} \rangle\rangle$ for $\langle k_{i\tilde{j}} \rangle = \langle\langle k_{i\tilde{j}} \rangle\rangle$. Expanding $\langle w_{i\tilde{j}} \rangle$ around $\langle k_{i\tilde{j}} \rangle = \langle\langle k_{i\tilde{j}} \rangle\rangle$, we obtain

$$\langle w_{i\tilde{j}} \rangle = \langle\langle w_{i\tilde{j}} \rangle\rangle + \frac{\partial \langle w_{i\tilde{j}} \rangle}{\partial \langle k_{i\tilde{j}} \rangle} (\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle) + \frac{\partial^2 \langle w_{i\tilde{j}} \rangle}{\partial \langle k_{i\tilde{j}} \rangle^2} (\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle)^2 + \dots \quad (\text{S9})$$

Dividing both sides by $\langle\langle w_{i\tilde{j}} \rangle\rangle$, we obtain

$$\frac{\langle w_{i\tilde{j}} \rangle}{\langle\langle w_{i\tilde{j}} \rangle\rangle} = 1 + \frac{1}{\langle\langle w_{i\tilde{j}} \rangle\rangle} \frac{\partial \langle w_{i\tilde{j}} \rangle}{\partial \langle k_{i\tilde{j}} \rangle} (\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle) + \frac{1}{\langle\langle w_{i\tilde{j}} \rangle\rangle} \frac{\partial^2 \langle w_{i\tilde{j}} \rangle}{\partial \langle k_{i\tilde{j}} \rangle^2} (\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle)^2 + \dots \quad (\text{S10})$$

By the definition of selection strength (see the main text under “Model”),

$$\left. \frac{1}{\langle\langle w_{i\tilde{j}} \rangle\rangle} \frac{\partial \langle w_{i\tilde{j}} \rangle}{\partial \langle k_{i\tilde{j}} \rangle} \right|_{\langle k_{i\tilde{j}} \rangle = \langle\langle k_{i\tilde{j}} \rangle\rangle} = \left(\frac{1}{\langle w_{i\tilde{j}} \rangle} \frac{\partial \langle w_{i\tilde{j}} \rangle}{\partial \langle k_{i\tilde{j}} \rangle} \right) \Big|_{\langle k_{i\tilde{j}} \rangle = \langle\langle k_{i\tilde{j}} \rangle\rangle} \quad (\text{S11})$$

$:= s_a$

By mathematical induction, it can be shown that

$$\frac{1}{\langle w_{i\tilde{j}} \rangle} \frac{\partial^{l+1} \langle w_{i\tilde{j}} \rangle}{\partial \langle k_{i\tilde{j}} \rangle^{l+1}} = \left(\frac{\partial}{\partial \langle k_{i\tilde{j}} \rangle} + \frac{1}{\langle w_{i\tilde{j}} \rangle} \frac{\partial \langle w_{i\tilde{j}} \rangle}{\partial \langle k_{i\tilde{j}} \rangle} \right) \frac{1}{\langle w_{i\tilde{j}} \rangle} \frac{\partial^l \langle w_{i\tilde{j}} \rangle}{\partial \langle k_{i\tilde{j}} \rangle^l} \quad (\text{S12})$$

for $l \in \{1, 2, 3, \dots\}$. Since it is assumed that $\partial s_a / \partial \langle k_{i\tilde{j}} \rangle = 0$ (see the main text under “Model”), the above equation implies that

$$\left(\frac{1}{\langle w_{i\tilde{j}} \rangle} \frac{\partial^l \langle w_{i\tilde{j}} \rangle}{\partial \langle k_{i\tilde{j}} \rangle^l} \right) \Big|_{\langle k_{i\tilde{j}} \rangle = \langle\langle k_{i\tilde{j}} \rangle\rangle} = s_a^l. \quad (\text{S13})$$

Given the above equation, Eq. (S10) implies that

$$\frac{\langle w_{i\tilde{j}} \rangle}{\langle\langle w_{i\tilde{j}} \rangle\rangle} = 1 + s_a (\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle) + O(s_a^2). \quad (\text{S14})$$

Therefore,

$$\begin{aligned} \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \text{cov}_{i\tilde{j}} [\langle w_{i\tilde{j}} \rangle, \langle k_{i\tilde{j}} \rangle] &= \text{cov}_{i\tilde{j}} [1 + s_a (\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle) + O(s_a^2), \langle k_{i\tilde{j}} \rangle] \\ &= s_a v_a + O(s_a^2) \end{aligned} \quad (\text{S15})$$

Second, we obtain the second term of Eq. (4), which stems from the second term of Eq. (3). Using the same method as above, we can show that

$$\langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}} [w_{ij}, k_{ij}] = -s_w v_{wi} + O(s_w^2). \quad (\text{S16})$$

We assume that $\langle w_{i\tilde{j}} \rangle$ and v_{wi} are statistically uncorrelated as i varies (this is equivalent to assuming that $\langle k_{i\tilde{j}} \rangle$ and v_{wi} are uncorrelated). Under this assumption,

$$\begin{aligned} \text{ave}_{i\tilde{j}} [\text{cov}_{i\tilde{j}} [w_{ij}, k_{ij}]] &= \text{ave}_{i\tilde{j}} [-\langle w_{i\tilde{j}} \rangle s_w v_{wi} + O(s_w^2)] \\ &= -\text{ave}_{i\tilde{j}} [\langle w_{i\tilde{j}} \rangle] s_w \text{ave}_{i\tilde{j}} [v_{wi}] + O(s_w^2) \\ &= -\langle\langle w_{i\tilde{j}} \rangle\rangle s_w v_w + O(s_w^2). \end{aligned} \quad (\text{S17})$$

Substituting Eqs. (S15) and (S17) into Eq. (3), we obtain Eq. (4).

3 Derivation of Eq. (8)

In this section, we derive Eq. (8), which is redisplayed below:

$$\begin{aligned}\mathbb{E}[v'_w] &= (1 - \beta N^{-1}) [v_w + m\sigma - s_w c_w + O(s_w^2)] \\ \mathbb{E}[v'_a] &= (1 - M^{-1}) [v_a + s_a c_a + O((s_w + s_a)^2)] \\ &\quad + (\beta N^{-1} - M^{-1}) [v_w + m\sigma - s_w c_w + O(s_w^2)],\end{aligned}$$

3.1 Calculation of $\mathbb{E}[v'_w]$

To calculate $\mathbb{E}[v'_w]$, we introduce the following symbols. Let n'_i be the number of replicators in collective i after one iteration of the Wright-Fisher process. Note that n'_i is a random variable and can be expressed as

$$n'_i = \sum_{l=1}^M \delta_{I_l i} \quad (\text{S18})$$

where $\delta_{I_l i}$ is the Kronecker delta (i.e., $\delta_{I_l i} = 1$ if $I_l = i$, and $\delta_{I_l i} = 0$ otherwise). Moreover, let $\langle k_{iJ_l} \rangle$ and $\langle \epsilon_{iJ_l} \rangle$ be the sample mean of k_{iJ_l} and ϵ_{iJ_l} within collective i :

$$\begin{aligned}\langle k_{iJ_l} \rangle &:= \frac{1}{n'_i} \sum_{l=1}^{n'_i} k_{iJ_l} \\ \langle \epsilon_{iJ_l} \rangle &:= \frac{1}{n'_i} \sum_{l=1}^{n'_i} \epsilon_{iJ_l},\end{aligned} \quad (\text{S19})$$

which are defined to be zero when $n'_i = 0$. The probability that $J_l = j$ given $I_l = i$ is

$$\begin{aligned}P(J = j | I = i) &= \frac{P(I = i, J = i)}{P(I = i)} \\ &= \frac{P(I = i, J = i)}{\sum_{j=1}^{n_i} P(I = i, J = i)} \\ &= \frac{\frac{w_{ij}}{M \langle w_{ij} \rangle}}{\sum_{j=1}^{n_i} \frac{w_{ij}}{M \langle w_{ij} \rangle}} \\ &= \frac{\frac{w_{ij}}{M \langle w_{ij} \rangle}}{\frac{n_i \langle w_{ij} \rangle}{M \langle w_{ij} \rangle}} \\ &= \frac{w_{ij}}{n_i \langle w_{ij} \rangle},\end{aligned} \quad (\text{S20})$$

where Eq. (S3) is used.

Using the symbols defined above, we can express the sample variance of k_{ij} within collective i in the next generation as

$$v'_{wi} := \frac{1}{n'_i} \sum_{l=1}^{n'_i} (k_{iJ_l} + \epsilon_{iJ_l} - \langle k_{iJ_l} \rangle - \langle \epsilon_{iJ_l} \rangle)^2, \quad (\text{S21})$$

which is defined to be zero when $n'_i = 0$. Using v'_{wi} , we can express $\mathbb{E}[v'_w]$ as follows:

$$\mathbb{E}[v'_w] = \mathbb{E}\left[\frac{1}{M} \sum_{i=1}^L n'_i v'_{wi}\right]. \quad (\text{S22})$$

In the last equation, we can separate k_{iJ_l} and ϵ_{iJ_l} as follows:

$$\begin{aligned} \mathbb{E}[v'_w] &= \mathbb{E}\left[\frac{1}{M} \sum_{i=1}^L n'_i \frac{1}{n'_i} \sum_{l=1}^{n'_i} (k_{iJ_l} + \epsilon_{iJ_l} - \langle k_{iJ_l} \rangle - \langle \epsilon_{iJ_l} \rangle)^2\right] \\ &= \mathbb{E}\left[\frac{1}{M} \sum_{i=1}^L \sum_{l=1}^{n'_i} \left\{ (k_{iJ_l} - \langle k_{iJ_l} \rangle)^2 + (\epsilon_{iJ_l} - \langle \epsilon_{iJ_l} \rangle)^2 + 2 (k_{iJ_l} - \langle k_{iJ_l} \rangle) (\epsilon_{iJ_l} - \langle \epsilon_{iJ_l} \rangle) \right\}\right] \\ &= \mathbb{E}\left[\frac{1}{M} \sum_{i=1}^L \sum_{l=1}^{n'_i} (k_{iJ_l} - \langle k_{iJ_l} \rangle)^2\right] + \mathbb{E}\left[\frac{1}{M} \sum_{i=1}^L \sum_{l=1}^{n'_i} (\epsilon_{iJ_l} - \langle \epsilon_{iJ_l} \rangle)^2\right] \\ &\quad + 2\mathbb{E}\left[\frac{1}{M} \sum_{i=1}^L \sum_{l=1}^{n'_i} (k_{iJ_l} - \langle k_{iJ_l} \rangle) (\epsilon_{iJ_l} - \langle \epsilon_{iJ_l} \rangle)\right]. \end{aligned} \quad (\text{S23})$$

The last term of the final line of Eq. (S23) can be shown to be zero, as follows. With the Kronecker delta δ_{Ii} , this term can be calculated as

$$\begin{aligned} &\mathbb{E}\left[\frac{1}{M} \sum_{i=1}^L \sum_{l=1}^{n'_i} (k_{iJ_l} - \langle k_{iJ_l} \rangle) (\epsilon_{iJ_l} - \langle \epsilon_{iJ_l} \rangle)\right] \\ &= \mathbb{E}\left[\frac{1}{M} \sum_{i=1}^L \sum_{l=1}^M \delta_{Ii} (k_{iJ_l} - \langle k_{iJ_l} \rangle) (\epsilon_{iJ_l} - \langle \epsilon_{iJ_l} \rangle)\right] \\ &= \frac{1}{M} \sum_{i=1}^L \sum_{l=1}^M \mathbb{E}[\delta_{Ii} (k_{iJ_l} - \langle k_{iJ_l} \rangle)] \mathbb{E}[(\epsilon_{iJ_l} - \langle \epsilon_{iJ_l} \rangle)] \\ &= 0, \end{aligned} \quad (\text{S24})$$

where we used the fact that k_{iJ_l} and ϵ_{iJ_l} are independent of each other. Therefore,

$$\mathbb{E}[v'_w] = \mathbb{E}\left[\frac{1}{M} \sum_{i=1}^L \sum_{l=1}^{n'_i} (k_{iJ_l} - \langle k_{iJ_l} \rangle)^2\right] + \mathbb{E}\left[\frac{1}{M} \sum_{i=1}^L \sum_{l=1}^{n'_i} (\epsilon_{iJ_l} - \langle \epsilon_{iJ_l} \rangle)^2\right]. \quad (\text{S25})$$

To calculate the first term of Eq. (S25), we define within-collective conditional expectation as follows:

$$\mathbb{E}_{J|I=i}[x_{IJ}] := \sum_{j=1}^{n_i} P(J = j | I = i) x_{ij}. \quad (\text{S26})$$

Using Eq. (S26), we can transform the first term of Eq. (S25) as follows:

$$\begin{aligned}
& \mathbb{E} \left[\frac{1}{M} \sum_{i=1}^L \sum_{l=1}^{n'_i} (k_{iJ_l} - \langle k_{iJ_l} \rangle)^2 \right] \\
&= \mathbb{E} \left[\frac{1}{M} \sum_{i=1}^L \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}] + \mathbb{E}_{J|I=i}[k_{IJ}] - \langle k_{iJ_l} \rangle)^2 \right] \\
&= \mathbb{E} \left[\frac{1}{M} \sum_{i=1}^L \left\{ \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}])^2 + \sum_{l=1}^{n'_i} (\mathbb{E}_{J|I=i}[k_{IJ}] - \langle k_{iJ_l} \rangle)^2 \right. \right. \\
&\quad \left. \left. + 2 \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}]) (\mathbb{E}_{J|I=i}[k_{IJ}] - \langle k_{iJ_l} \rangle) \right\} \right] \\
&= \mathbb{E} \left[\frac{1}{M} \sum_{i=1}^L \left\{ \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}])^2 + n'_i (\mathbb{E}_{J|I=i}[k_{IJ}] - \langle k_{iJ_l} \rangle)^2 \right. \right. \\
&\quad \left. \left. + 2 (\mathbb{E}_{J|I=i}[k_{IJ}] - \langle k_{iJ_l} \rangle) \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}]) \right\} \right] \tag{S27} \\
&= \mathbb{E} \left[\frac{1}{M} \sum_{i=1}^L \left\{ \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}])^2 + n'_i (\mathbb{E}_{J|I=i}[k_{IJ}] - \langle k_{iJ_l} \rangle)^2 \right. \right. \\
&\quad \left. \left. + 2 (\mathbb{E}_{J|I=i}[k_{IJ}] - \langle k_{iJ_l} \rangle) n'_i (\langle k_{iJ_l} \rangle - \mathbb{E}_{J|I=i}[k_{IJ}]) \right\} \right] \\
&= \mathbb{E} \left[\frac{1}{M} \sum_{i=1}^L \left\{ \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}])^2 - n'_i (\langle k_{iJ_l} \rangle - \mathbb{E}_{J|I=i}[k_{IJ}])^2 \right\} \right] \\
&= \frac{1}{M} \sum_{i=1}^L \mathbb{E} \left[\sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}])^2 \right] - \frac{1}{M} \sum_{i=1}^L \mathbb{E} \left[n'_i (\langle k_{iJ_l} \rangle - \mathbb{E}_{J|I=i}[k_{IJ}])^2 \right].
\end{aligned}$$

The first term in the last line of Eq. (S27) is calculated as follows:

$$\begin{aligned}
& \frac{1}{M} \sum_{i=1}^L \mathbb{E} \left[\sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}])^2 \right] \\
&= \frac{1}{M} \sum_{i=1}^L \mathbb{E} \left[\sum_{l=1}^{n'_i} (k_{iJ_l}^2 - 2\mathbb{E}_{J|I=i}[k_{IJ}]k_{iJ_l} + \mathbb{E}_{J|I=i}[k_{IJ}]^2) \right] \\
&= \frac{1}{M} \sum_{i=1}^L \mathbb{E} \left[\sum_{l=1}^{n'_i} k_{iJ_l}^2 - 2\mathbb{E}_{J|I=i}[k_{IJ}] \sum_{l=1}^{n'_i} k_{iJ_l} + n'_i \mathbb{E}_{J|I=i}[k_{IJ}]^2 \right] \\
&= \frac{1}{M} \sum_{i=1}^L \left\{ \mathbb{E} \left[\sum_{l=1}^{n'_i} k_{iJ_l}^2 \right] - 2\mathbb{E}_{J|I=i}[k_{IJ}] \mathbb{E} \left[\sum_{l=1}^{n'_i} k_{iJ_l} \right] + \mathbb{E}[n'_i] \mathbb{E}_{J|I=i}[k_{IJ}]^2 \right\}. \tag{S28}
\end{aligned}$$

Using the Kronecker delta δ_{Ii} , we can show that

$$\begin{aligned}
\mathbb{E}\left[\sum_{l=1}^{n'_i} k_{iJ_l}\right] &= \mathbb{E}\left[\sum_{l=1}^M \delta_{I_l i} k_{I_l J_l}\right] \\
&= \sum_{l=1}^M \mathbb{E}[\delta_{I_l i} k_{I_l J_l}] \\
&= M \mathbb{E}[\delta_{Ii} k_{IJ}] \\
&= M \sum_{j=1}^{n_i} P(I = i, J = j) k_{ij} \\
&= M P(I = i) \sum_{j=1}^{n_i} P(J = j | I = i) k_{ij} \\
&= M \frac{n_i \langle w_{i\tilde{j}} \rangle}{M \langle\langle w_{i\tilde{j}} \rangle\rangle} \mathbb{E}_{J|I=i}[k_{IJ}] \\
&= \frac{n_i \langle w_{i\tilde{j}} \rangle}{\langle\langle w_{i\tilde{j}} \rangle\rangle} \mathbb{E}_{J|I=i}[k_{IJ}].
\end{aligned} \tag{S29}$$

Likewise, we can show that

$$\begin{aligned}
\mathbb{E}\left[\sum_{l=1}^{n'_i} k_{iJ_l}^2\right] &= \mathbb{E}\left[\sum_{l=1}^M \delta_{I_l i} k_{I_l J_l}^2\right] \\
&= \sum_{l=1}^M \mathbb{E}[\delta_{I_l i} k_{I_l J_l}^2] \\
&= M \mathbb{E}[\delta_{Ii} k_{IJ}^2] \\
&= M \sum_{j=1}^{n_i} P(I = i, J = j) k_{ij}^2 \\
&= M P(I = i) \sum_{j=1}^{n_i} P(J = j | I = i) k_{ij}^2 \\
&= \frac{n_i \langle w_{i\tilde{j}} \rangle}{\langle\langle w_{i\tilde{j}} \rangle\rangle} \mathbb{E}_{J|I=i}[k_{IJ}^2].
\end{aligned} \tag{S30}$$

Also, we can show that

$$\begin{aligned}
\mathbb{E}[n'_i] &= \mathbb{E}\left[\sum_{l=1}^M \delta_{Ili}\right] \\
&= \sum_{l=1}^M \mathbb{E}[\delta_{Ili}] \\
&= M\mathbb{E}[\delta_{Ii}] \\
&= M \sum_{j=1}^{n_i} P(I=i, J=j) \\
&= MP(I=i) \\
&= \frac{n_i \langle w_{i\tilde{j}} \rangle}{\langle\langle w_{i\tilde{j}} \rangle\rangle}.
\end{aligned} \tag{S31}$$

Using the above results, we can transform the last line of Eq. (S28) as follows:

$$\begin{aligned}
&\frac{1}{M} \sum_{i=1}^L \left\{ \mathbb{E}\left[\sum_{l=1}^{n'_i} k_{iJ_l}^2\right] - 2\mathbb{E}_{J|I=i}[k_{IJ}]\mathbb{E}\left[\sum_{l=1}^{n'_i} k_{iJ_l}\right] + \mathbb{E}[n'_i]\mathbb{E}_{J|I=i}[k_{IJ}]^2 \right\} \\
&= \frac{1}{M} \sum_{i=1}^L \left\{ \frac{n_i \langle w_{i\tilde{j}} \rangle}{\langle\langle w_{i\tilde{j}} \rangle\rangle} \mathbb{E}_{J|I=i}[k_{IJ}^2] - 2\mathbb{E}_{J|I=i}[k_{IJ}] \frac{n_i \langle w_{i\tilde{j}} \rangle}{\langle\langle w_{i\tilde{j}} \rangle\rangle} \mathbb{E}_{J|I=i}[k_{IJ}] + \frac{n_i \langle w_{i\tilde{j}} \rangle}{\langle\langle w_{i\tilde{j}} \rangle\rangle} \mathbb{E}_{J|I=i}[k_{IJ}]^2 \right\} \\
&= \frac{1}{M} \sum_{i=1}^L \left\{ \frac{n_i \langle w_{i\tilde{j}} \rangle}{\langle\langle w_{i\tilde{j}} \rangle\rangle} \mathbb{E}_{J|I=i}[k_{IJ}^2] - \frac{n_i \langle w_{i\tilde{j}} \rangle}{\langle\langle w_{i\tilde{j}} \rangle\rangle} \mathbb{E}_{J|I=i}[k_{IJ}]^2 \right\} \\
&= \frac{1}{M} \sum_{i=1}^L \frac{n_i \langle w_{i\tilde{j}} \rangle}{\langle\langle w_{i\tilde{j}} \rangle\rangle} \left\{ \mathbb{E}_{J|I=i}[k_{IJ}^2] - \mathbb{E}_{J|I=i}[k_{IJ}]^2 \right\}.
\end{aligned} \tag{S32}$$

The conditional expectation in the last line of Eq. (S32) can be calculated as follows:

$$\begin{aligned}
&\mathbb{E}_{J|I=i}[k_{IJ}^2] - \mathbb{E}_{J|I=i}[k_{IJ}]^2 \\
&= \sum_{j=1}^{n_i} P(J=j|i=i) k_{ij}^2 - \left(\sum_{j=1}^{n_i} P(J=j|i=i) k_{ij} \right)^2 \\
&= \sum_{j=1}^{n_i} \frac{w_{ij}}{n_i \langle w_{i\tilde{j}} \rangle} k_{ij}^2 - \left(\sum_{j=1}^{n_i} \frac{w_{ij}}{n_i \langle w_{i\tilde{j}} \rangle} k_{ij} \right)^2 \\
&= \langle w_{i\tilde{j}} \rangle^{-1} \frac{1}{n_i} \sum_{j=1}^{n_i} w_{ij} k_{ij}^2 - \left(\langle w_{i\tilde{j}} \rangle^{-1} \frac{1}{n_i} \sum_{j=1}^{n_i} w_{ij} k_{ij} \right)^2 \\
&= \langle w_{i\tilde{j}} \rangle^{-1} \left(\text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}^2] + \langle w_{i\tilde{j}} \rangle \langle k_{ij}^2 \rangle \right) - \left\{ \langle w_{i\tilde{j}} \rangle^{-1} (\text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}] + \langle w_{i\tilde{j}} \rangle \langle k_{ij} \rangle) \right\}^2 \\
&= \langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}^2] + \langle k_{ij}^2 \rangle - \left\{ \langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}] + \langle k_{ij} \rangle \right\}^2 \\
&= v_{\mathbf{wi}} + \langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}^2] - \left\{ \langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}] \right\}^2 - 2\langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}] \langle k_{ij} \rangle \\
&= v_{\mathbf{wi}} + \langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}}[w_{ij}, (k_{ij} - \langle k_{i\tilde{j}} \rangle)^2] - \left\{ \langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}] \right\}^2,
\end{aligned} \tag{S33}$$

where we used the fact that $\langle k_{i\tilde{j}}^2 \rangle - \langle k_{i\tilde{j}} \rangle^2 = v_{wi}$. The last line of Eq. (S33) can be interpreted as the expected variance of k_{ij} within collective i after one iteration of the Wright-Fisher process excluding the effect of random sampling. Thus, let us introduce the following symbol:

$$\begin{aligned} \Delta_s v_{wi} &:= (\mathbb{E}_{J|I=i}[k_{IJ}^2] - \mathbb{E}_{J|I=i}[k_{IJ}]^2) - v_{wi} \\ &= \langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}}[w_{ij}, (k_{ij} - \langle k_{i\tilde{j}} \rangle)^2] - \{ \langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}] \}^2, \end{aligned} \quad (\text{S34})$$

which denotes the expected change of the variance of k_{ij} within collective i due to within-collective selection.

Combining Eqs. (S28), (S32), (S33), and (S34), we can transform the first term in the last line of Eq. (S27) as follows

$$\frac{1}{M} \sum_{i=1}^L \mathbb{E} \left[\sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}])^2 \right] = \frac{1}{M} \sum_{i=1}^L \frac{n_i \langle w_{i\tilde{j}} \rangle}{\langle\langle w_{i\tilde{j}} \rangle\rangle} (v_{wi} + \Delta_s v_{wi}) \quad (\text{S35})$$

Next, we calculate the second term in the last line of Eq. (S27) as follows:

$$\begin{aligned} & \frac{1}{M} \sum_{i=1}^L \mathbb{E} \left[n'_i (\langle k_{iJ_l} \rangle - \mathbb{E}_{J|I=i}[k_{IJ}])^2 \right] \\ &= \frac{1}{M} \sum_{i=1}^L \mathbb{E} \left[n'_i \left(\frac{1}{n'_i} \sum_{l=1}^{n'_i} k_{iJ_l} - \frac{1}{n'_i} \sum_{l=1}^{n'_i} \mathbb{E}_{J|I=i}[k_{IJ}] \right)^2 \right] \\ &= \frac{1}{M} \sum_{i=1}^L \mathbb{E} \left[n'_i \left\{ \frac{1}{n'_i} \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}]) \right\}^2 \right] \\ &= \frac{1}{M} \sum_{i=1}^L \mathbb{E} \left[\frac{1}{n'_i} \left\{ \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}]) \right\}^2 \right] \\ &= \frac{1}{M} \sum_{i=1}^L \mathbb{E} \left[\frac{1}{n'_i} \left\{ \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}])^2 \right. \right. \\ & \quad \left. \left. + 2 \sum_{m \neq l} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}]) (k_{iJ_m} - \mathbb{E}_{J|I=i}[k_{IJ}]) \right\} \right] \\ &= \frac{1}{M} \sum_{i=1}^L \mathbb{E} \left[\frac{1}{n'_i} \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}])^2 \right], \end{aligned} \quad (\text{S36})$$

where we used the fact that k_{iJ_l} and k_{iJ_m} are independent of each other for $l \neq m$ in the final step. Since n'_i can be zero, the last line of Eq. (S36) needs to be interpreted as follows:

$$\frac{1}{n'_i} \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}])^2 = \begin{cases} 0, & \text{if } n'_i = 0 \\ \frac{1}{n'_i} \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}])^2, & \text{if } n'_i > 0. \end{cases} \quad (\text{S37})$$

Thus,

$$\mathbb{E} \left[\frac{1}{n'_i} \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}])^2 \right] = \begin{cases} 0, & \text{if } n'_i = 0 \\ \mathbb{E}_{J|I=i}[k_{IJ}^2] - \mathbb{E}_{J|I=i}[k_{IJ}]^2, & \text{if } n'_i > 0, \end{cases} \quad (\text{S38})$$

where the case for $n'_i > 0$ follows from a calculation similar to Eqs. (S28) and (S32).

To calculate the last line of Eq. (S36), we separate the case where $n'_i > 0 \forall i \in \{1, 2, 3, \dots, L\}$ and the case where $n'_i = 0$ for some $i \in \{1, 2, 3, \dots, L\}$. Let S be a proper subset of $\{1, 2, 3, \dots, L\}$, and $P(S)$ be the probability that $n'_i = 0 \forall i \in S$ and $n'_i > 0 \forall i \notin S$. The value of $P(S)$ can be estimated as follows:

$$\begin{aligned} P(S) &\leq \left[1 - \sum_{i \in S} P(I = i) \right]^M \\ &= \left[1 - \sum_{i \in S} \frac{n_i \langle w_{i\tilde{j}} \rangle}{M \langle\langle w_{i\tilde{j}} \rangle\rangle} \right]^M \\ &\approx \exp \left(- \sum_{i \in S} \frac{n_i \langle w_{i\tilde{j}} \rangle}{\langle\langle w_{i\tilde{j}} \rangle\rangle} \right), \end{aligned} \tag{S39}$$

where the RHS of the first inequality is the probability that $n'_i = 0 \forall i \in S$ and $n'_i \geq 0 \forall i \notin S$. Using these symbols and Eq. (S34), we can express the last line of Eq. (S36) as follows:

$$\begin{aligned} &\frac{1}{M} \sum_{i=1}^L \mathbb{E} \left[\frac{1}{n'_i} \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}])^2 \right], \\ &= \frac{1}{M} \sum_S P(S) \sum_{i \notin S} (\mathbb{E}_{J|I=i}[k_{IJ}^2] - \mathbb{E}_{J|I=i}[k_{IJ}]^2) \\ &= \frac{1}{M} \sum_S P(S) \sum_{i \notin S} (v_{wi} + \Delta_s v_{wi}) \\ &= \frac{1}{M} \sum_S P(S) \left[\sum_{i=1}^L (v_{wi} + \Delta_s v_{wi}) - \sum_{i \in S} (v_{wi} + \Delta_s v_{wi}) \right] \\ &= \frac{1}{M} \sum_{i=1}^L (v_{wi} + \Delta_s v_{wi}) - \frac{1}{M} \sum_{S \neq \emptyset} P(S) \sum_{i \in S} (v_{wi} + \Delta_s v_{wi}) \end{aligned} \tag{S40}$$

where \sum_S is a summation over all possible $S \subset \{1, 2, \dots, L\}$, and $\sum_{S \neq \emptyset}$ is the same summation excluding the case where S is empty.

We assume that the second term of the last line of Eq. (S40) is negligible for the following reasons. If $n_i \gg 1$ for some $i \in S$, then Eq. (S39) implies that $P(S) \approx 0$. Contrariwise, if the statement that $n_i \gg 1$ is false for all $i \in S$, then the value of $\sum_{i \in S} (v_{wi} + \Delta_s v_{wi})$ is likely to be small because n_i is small. Under this assumption, Eq. (S40) implies that

$$\frac{1}{M} \sum_{i=1}^L \mathbb{E} \left[\frac{1}{n'_i} \sum_{l=1}^{n'_i} (k_{iJ_l} - \mathbb{E}_{J|I=i}[k_{IJ}])^2 \right] \approx \frac{1}{M} \sum_{i=1}^L (v_{wi} + \Delta_s v_{wi}). \tag{S41}$$

The second term of Eq. (S25) is calculated as follows:

$$\begin{aligned}
& \mathbb{E} \left[\frac{1}{M} \sum_{i=1}^L \sum_{l=1}^{n'_i} (\epsilon_{iJ_l} - \langle \epsilon_{iJ_l} \rangle)^2 \right] \\
&= \mathbb{E} \left[\frac{1}{M} \sum_{i=1}^L \sum_{l=1}^{n'_i} (\epsilon_{iJ_l}^2 - 2\langle \epsilon_{iJ_l} \rangle \epsilon_{iJ_l} + \langle \epsilon_{iJ_l} \rangle^2) \right] \\
&= \mathbb{E} \left[\frac{1}{M} \sum_{i=1}^L \left(\sum_{l=1}^{n'_i} \epsilon_{iJ_l}^2 - 2\langle \epsilon_{iJ_l} \rangle \sum_{l=1}^{n'_i} \epsilon_{iJ_l} + \langle \epsilon_{iJ_l} \rangle^2 \sum_{l=1}^{n'_i} 1 \right) \right] \\
&= \mathbb{E} \left[\frac{1}{M} \sum_{i=1}^L \left(\sum_{l=1}^{n'_i} \epsilon_{iJ_l}^2 - 2n'_i \langle \epsilon_{iJ_l} \rangle + n'_i \langle \epsilon_{iJ_l} \rangle^2 \right) \right] \\
&= \mathbb{E} \left[\frac{1}{M} \sum_{i=1}^L \left(\sum_{l=1}^{n'_i} \epsilon_{iJ_l}^2 - n'_i \langle \epsilon_{iJ_l} \rangle^2 \right) \right] \\
&= \mathbb{E} \left[\frac{1}{M} \sum_{l=1}^M \epsilon_{I_l J_l}^2 - \frac{1}{M} \sum_{i=1}^L n'_i \langle \epsilon_{iJ_l} \rangle^2 \right] \\
&= \mathbb{E} \left[\frac{1}{M} \sum_{l=1}^M \epsilon_{I_l J_l}^2 - \frac{1}{M} \sum_{i=1}^L n'_i \left(\frac{1}{n'_i} \sum_{l=1}^{n'_i} \epsilon_{I_l J_l} \right)^2 \right] \\
&= m\sigma - \frac{1}{M} \mathbb{E} \left[\sum_{i=1}^L \frac{1}{n'_i} \left(\sum_{l=1}^{n'_i} \epsilon_{I_l J_l} \right)^2 \right] \\
&= m\sigma - \frac{1}{M} \mathbb{E} \left[\sum_{i=1}^L \frac{1}{n'_i} \left(\sum_{l=1}^{n'_i} \epsilon_{I_l J_l}^2 + 2 \sum_{l \neq m} \epsilon_{I_l J_l} \epsilon_{I_m J_m} \right) \right] \\
&= m\sigma - \frac{1}{M} \mathbb{E} \left[\sum_{i=1}^L \frac{1}{n'_i} \sum_{l=1}^{n'_i} \epsilon_{I_l J_l}^2 \right] \\
&= m\sigma - \frac{1}{M} \sum_S P(S) \sum_{i \notin S} \mathbb{E}[\epsilon_{I_l J_l}^2] \\
&= m\sigma - \frac{1}{M} \sum_S P(S) \left(\sum_{i=1}^L m\sigma - \sum_{i \in S} m\sigma \right) \\
&= m\sigma - \frac{L}{M} m\sigma + \frac{1}{M} \sum_S P(S) \sum_{i \in S} m\sigma \\
&= \left(1 - \frac{L}{M} + \frac{1}{M} \sum_S P(S) |S| \right) m\sigma, \\
&\approx \left(1 - \frac{L}{M} \right) m\sigma,
\end{aligned} \tag{S42}$$

where $|S|$ is the number of elements in S , and we have assumed that $\sum_S P(S) |S| \ll L$ in the last step.

Combining Eqs. (S25), (S27), (S35), (S36), (S41), and (S42), we obtain

$$\begin{aligned}
\mathbb{E}[v'_w] &\approx \frac{1}{M} \sum_{i=1}^L \frac{n_i \langle w_{i\tilde{j}} \rangle}{\langle\langle w_{i\tilde{j}} \rangle\rangle} \{v_{wi} + \Delta_s v_{wi}\} \\
&\quad - \frac{1}{M} \sum_{i=1}^L \{v_{wi} + \Delta_s v_{wi}\} + \left(1 - \frac{L}{M}\right) m\sigma \\
&= \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \text{cov}_{i\tilde{j}} [\langle w_{i\tilde{j}} \rangle, v_{wi} + \Delta_s v_{wi}] + M^{-1} \sum_{i=1}^L n_i (v_{wi} + \Delta_s v_{wi}) \\
&\quad - M^{-1} \sum_{i=1}^L n_i n_i^{-1} \{v_{wi} + \Delta_s v_{wi}\} + M^{-1} \sum_{i=1}^L n_i (1 - n_i^{-1}) m\sigma. \\
&= \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \text{cov}_{i\tilde{j}} [\langle w_{i\tilde{j}} \rangle, v_{wi} + \Delta_s v_{wi}] + M^{-1} \sum_{i=1}^L n_i (1 - n_i^{-1}) (v_{wi} + \Delta_s v_{wi} + m\sigma) \\
&= \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \text{cov}_{i\tilde{j}} [\langle w_{i\tilde{j}} \rangle, v_{wi} + \Delta_s v_{wi}] + \text{ave}_{i\tilde{j}} [(1 - n_i^{-1}) (v_{wi} + \Delta_s v_{wi} + m\sigma)] \\
&= \text{ave}_{i\tilde{j}} [(1 - n_i^{-1}) (v_{wi} + \Delta_s v_{wi} + m\sigma)],
\end{aligned} \tag{S43}$$

where we have assumed in the last step that $\langle w_{i\tilde{j}} \rangle$ and $v_{wi} + \Delta_s v_{wi}$ are statistically uncorrelated as i varies, as we have assumed in Eq. (S17).

To enable the further calculation of Eq. (S43), we assume that

$$n_i = \beta N^{-1}. \tag{S44}$$

Under this assumption, we can transform Eq. (S43) as follows

$$\begin{aligned}
\mathbb{E}[v'_w] &\approx \text{ave}_{i\tilde{j}} [(1 - n_i^{-1}) (v_{wi} + m\sigma + \Delta_s v_{wi})] \\
&\approx (1 - \beta N^{-1}) \text{ave}_{i\tilde{j}} \left[v_{wi} + m\sigma + \langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}} [w_{ij}, (k_{ij} - \langle k_{i\tilde{j}} \rangle)^2] \right. \\
&\quad \left. - \{ \langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}} [w_{ij}, k_{ij}] \}^2 \right]. \\
&= (1 - \beta N^{-1}) \left(v_w + m\sigma + \text{ave}_{i\tilde{j}} [\langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}} [w_{ij}, (k_{ij} - \langle k_{i\tilde{j}} \rangle)^2]] + O(s_w^2) \right),
\end{aligned} \tag{S45}$$

where we used Eq. (S16) in the last step. Expanding w_{ij} as a Taylor series around $k_{ij} = \langle k_{i\tilde{j}} \rangle$, we can show that

$$\langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}} [w_{ij}, (k_{ij} - \langle k_{i\tilde{j}} \rangle)^2] = -s_w c_{wi} + O(s_w^2), \tag{S46}$$

where $c_{wi} := \langle (k_{i\tilde{j}} - \langle k_{i\tilde{j}} \rangle)^3 \rangle$, and that

$$\text{ave}_{i\tilde{j}} [\langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}} [w_{ij}, (k_{ij} - \langle k_{i\tilde{j}} \rangle)^2]] = -s_w c_w + O(s_w^2), \tag{S47}$$

where $c_w := \text{ave}_{i\tilde{j}} [c_{wi}]$. Substituting Eq. (S47) into Eq. (S45), we obtain

$$\mathbb{E}[v'_w] \approx (1 - \beta N^{-1}) (v_w + m\sigma - s_w c_w + O(s_w^2)). \tag{S48}$$

3.2 Calculation of $\mathbb{E}[v'_a]$

To calculate $\mathbb{E}[v'_a]$, we first calculate $\mathbb{E}[v'_t]$ and then use the fact that $\mathbb{E}[v'_a] = \mathbb{E}[v'_t] - \mathbb{E}[v'_w]$.

We can express $\mathbb{E}[v'_t]$ as follows:

$$\begin{aligned}\mathbb{E}[v'_t] &= \mathbb{E}\left[\frac{1}{M} \sum_{l=1}^M \left\{ k_{I_l J_l} + \epsilon_{i_{J_l}} - \frac{1}{M} \sum_{l=1}^M (k_{I_l J_l} + \epsilon_{i_{J_l}}) \right\}^2\right] \\ &= \left(1 - \frac{1}{M}\right) \mathbb{E}[(k_{IJ} - \mathbb{E}[k_{IJ}])^2] + \left(1 - \frac{1}{M}\right) m\sigma,\end{aligned}\tag{S49}$$

where we used the fact that $k_{I_l J_l}$ and $\epsilon_{i_{J_l}}$ are independent of each other, as we did in Eqs. (S23), (S24), and (S25).

The first term of the last line of Eq. (S49) is calculated as follows:

$$\begin{aligned}&\mathbb{E}[(k_{IJ} - \mathbb{E}[k_{IJ}])^2] \\ &= \mathbb{E}\left[\left(k_{ij} - \mathbb{E}_{J|I=i}[k_{IJ}] + \mathbb{E}_{J|I=i}[k_{IJ}] - \mathbb{E}[k_{IJ}]\right)^2\right] \\ &= \mathbb{E}\left[\left(k_{ij} - \mathbb{E}_{J|I=i}[k_{IJ}]\right)^2\right] + \mathbb{E}\left[\left(\mathbb{E}_{J|I=i}[k_{IJ}] - \mathbb{E}[k_{IJ}]\right)^2\right] \\ &\quad - 2\mathbb{E}\left[\left(k_{ij} - \mathbb{E}_{J|I=i}[k_{IJ}]\right)\left(\mathbb{E}_{J|I=i}[k_{IJ}] - \mathbb{E}[k_{IJ}]\right)\right] \\ &= \mathbb{E}\left[\left(k_{ij} - \mathbb{E}_{J|I=i}[k_{IJ}]\right)^2\right] + \mathbb{E}\left[\left(\mathbb{E}_{J|I=i}[k_{IJ}] - \mathbb{E}[k_{IJ}]\right)^2\right] \\ &\quad - 2 \sum_{i=1}^L \sum_{j=1}^{n_i} P(I=i, J=j) \left(\mathbb{E}_{J|I=i}[k_{IJ}] - \mathbb{E}[k_{IJ}]\right) \left(k_{ij} - \mathbb{E}_{J|I=i}[k_{IJ}]\right) \\ &= \mathbb{E}\left[\left(k_{ij} - \mathbb{E}_{J|I=i}[k_{IJ}]\right)^2\right] + \mathbb{E}\left[\left(\mathbb{E}_{J|I=i}[k_{IJ}] - \mathbb{E}[k_{IJ}]\right)^2\right] \\ &\quad - 2 \sum_{i=1}^L P(I=i) \left(\mathbb{E}_{J|I=i}[k_{IJ}] - \mathbb{E}[k_{IJ}]\right) \sum_{j=1}^{n_i} P(J=j|I=i) \left(k_{ij} - \mathbb{E}_{J|I=i}[k_{IJ}]\right) \\ &= \mathbb{E}\left[\left(k_{ij} - \mathbb{E}_{J|I=i}[k_{IJ}]\right)^2\right] + \mathbb{E}\left[\left(\mathbb{E}_{J|I=i}[k_{IJ}] - \mathbb{E}[k_{IJ}]\right)^2\right]\end{aligned}\tag{S50}$$

The first term of the last line of Eq. (S50) is calculated as follows:

$$\begin{aligned}
& \mathbb{E} \left[\left(k_{ij} - \mathbb{E}_{J|I=i}[k_{IJ}] \right)^2 \right] \\
&= \mathbb{E} \left[k_{ij}^2 - 2\mathbb{E}_{J|I=i}[k_{IJ}]k_{ij} + \mathbb{E}_{J|I=i}[k_{IJ}]^2 \right] \\
&= \sum_{i=1}^L \sum_{j=1}^{n_i} P(I=i, J=j) \left(k_{ij}^2 - 2\mathbb{E}_{J|I=i}[k_{IJ}]k_{ij} + \mathbb{E}_{J|I=i}[k_{IJ}]^2 \right) \\
&= \sum_{i=1}^L P(I=i) \left(\sum_{j=1}^{n_i} P(J=j|I=i) k_{ij}^2 \right. \\
&\quad \left. - 2\mathbb{E}_{J|I=i}[k_{IJ}] \sum_{j=1}^{n_i} P(J=j|I=i) k_{ij} + \mathbb{E}_{J|I=i}[k_{IJ}]^2 \right) \\
&= \sum_{i=1}^L P(I=i) \left(\mathbb{E}_{J|I=i}[k_{IJ}^2] - \mathbb{E}_{J|I=i}[k_{IJ}]^2 \right) \\
&= \sum_{i=1}^L \frac{n_i \langle w_{i\tilde{j}} \rangle}{M \langle\langle w_{i\tilde{j}} \rangle\rangle} \left(\mathbb{E}_{J|I=i}[k_{IJ}^2] - \mathbb{E}_{J|I=i}[k_{IJ}]^2 \right) \\
&= \sum_{i=1}^L \frac{n_i \langle w_{i\tilde{j}} \rangle}{M \langle\langle w_{i\tilde{j}} \rangle\rangle} (v_{wi} + \Delta_s v_{wi}), \\
&= \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \text{cov}_{\tilde{i}} [\langle w_{i\tilde{j}} \rangle, v_{wi} + \Delta_s v_{wi}] + M^{-1} \sum_{i=1}^L n_i (v_{wi} + \Delta_s v_{wi}) \\
&= \text{ave}_{\tilde{i}} [v_{wi} + \Delta_s v_{wi}],
\end{aligned} \tag{S51}$$

where we used Eq. (S34) and the assumption that $\langle w_{i\tilde{j}} \rangle$ and $v_{wi} + \Delta_s v_{wi}$ are statistically uncorrelated as i varies, which has already been made in Eq. (S17). Using Eq. (S16) and (S47), we can transform the last line of Eq. (S51) as follows:

$$\text{ave}_{\tilde{i}} [v_{wi} + \Delta_s v_{wi}] = v_w - s_w c_w + O(s_w^2). \tag{S52}$$

Therefore,

$$\mathbb{E} \left[\left(k_{ij} - \mathbb{E}_{J|I=i}[k_{IJ}] \right)^2 \right] = v_w - s_w c_w + O(s_w^2). \tag{S53}$$

The second term of the last line of Eq. (S50) is calculated as follows:

$$\begin{aligned}
& \mathbb{E} \left[\left(\mathbb{E}_{J|I=i}[k_{IJ}] - \mathbb{E}[k_{IJ}] \right)^2 \right] \\
&= \mathbb{E} \left[\mathbb{E}_{J|I=i}[k_{IJ}]^2 - 2\mathbb{E}[k_{IJ}]\mathbb{E}_{J|I=i}[k_{IJ}] + \mathbb{E}[k_{IJ}]^2 \right] \\
&= \sum_{i=1}^L \sum_{j=1}^{n_i} P(I=i, J=j) \left(\mathbb{E}_{J|I=i}[k_{IJ}]^2 - 2\mathbb{E}[k_{IJ}]\mathbb{E}_{J|I=i}[k_{IJ}] + \mathbb{E}[k_{IJ}]^2 \right) \\
&= \sum_{i=1}^L P(I=i) \mathbb{E}_{J|I=i}[k_{IJ}]^2 - 2\mathbb{E}[k_{IJ}] \sum_{i=1}^L P(I=i) \mathbb{E}_{J|I=i}[k_{IJ}] + \mathbb{E}[k_{IJ}]^2 \\
&= \sum_{i=1}^L P(I=i) \mathbb{E}_{J|I=i}[k_{IJ}]^2 - \mathbb{E}[k_{IJ}]^2
\end{aligned} \tag{S54}$$

The first term of the last line of Eq. (S54) is calculated as follows:

$$\begin{aligned}
& \sum_{i=1}^L P(I=i) \mathbb{E}_{J|I=i} [k_{IJ}]^2 \\
&= \sum_{i=1}^L P(I=i) \left[\sum_{j=1}^{n_i} P(J=j|I=i) k_{ij} \right]^2 \\
&= \sum_{i=1}^L P(I=i) \left[\sum_{j=1}^{n_i} \frac{w_{ij}}{n_i \langle w_{i\tilde{j}} \rangle} k_{ij} \right]^2 \\
&= \sum_{i=1}^L P(I=i) \left[\frac{1}{\langle w_{i\tilde{j}} \rangle n_i} \sum_{j=1}^{n_i} w_{ij} k_{ij} \right]^2 \\
&= \sum_{i=1}^L P(I=i) \left[\langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}] + \langle k_{i\tilde{j}} \rangle \right]^2 \tag{S55} \\
&= \sum_{i=1}^L \frac{n_i \langle w_{i\tilde{j}} \rangle}{M \langle\langle w_{i\tilde{j}} \rangle\rangle} \left[\langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}] + \langle k_{i\tilde{j}} \rangle \right]^2 \\
&= \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \sum_{i=1}^L \frac{n_i}{M} \langle w_{i\tilde{j}} \rangle \left[\langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}] + \langle k_{i\tilde{j}} \rangle \right]^2 \\
&= \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \text{cov}_{\tilde{i}} \left[\langle w_{i\tilde{j}} \rangle, \left\{ \langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}] + \langle k_{i\tilde{j}} \rangle \right\}^2 \right] \\
&\quad + \text{ave}_{\tilde{i}} \left[\left\{ \langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}] + \langle k_{i\tilde{j}} \rangle \right\}^2 \right] \\
&= \langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \text{cov}_{\tilde{i}} \left[\langle w_{i\tilde{j}} \rangle, (\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle)^2 \right] + \text{ave}_{\tilde{i}} \left[(\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle)^2 \right],
\end{aligned}$$

where we have used the following notation in the final step:

$$\Delta_s \langle k_{i\tilde{j}} \rangle := \langle w_{i\tilde{j}} \rangle^{-1} \text{cov}_{i\tilde{j}}[w_{ij}, k_{ij}]. \tag{S56}$$

The second term of the last line of Eq. (S54) is calculated as follows:

$$\begin{aligned}
\mathbb{E}[k_{IJ}]^2 &= \left(\sum_{i=1}^L \sum_{j=1}^{n_i} P(I=i, J=j) k_{ij} \right)^2 \\
&= \left(\sum_{i=1}^L P(I=i) \sum_{j=1}^{n_i} P(J=j|I=i) k_{ij} \right)^2 \\
&= \left(\sum_{i=1}^L P(I=i) \mathbb{E}_{J|I=i} [k_{IJ}] \right)^2 \\
&= \left[\sum_{i=1}^L \frac{n_i \langle w_{i\tilde{j}} \rangle}{M \langle\langle w_{i\tilde{j}} \rangle\rangle} \mathbb{E}_{J|I=i} [k_{IJ}] \right]^2. \tag{S57}
\end{aligned}$$

Doing the same calculation as in Eq. (S55), we can transform Eq. (S57) as follows:

$$\begin{aligned}
& \left[\sum_{i=1}^L \frac{n_i \langle w_{i\tilde{j}} \rangle}{M \langle \langle w_{i\tilde{j}} \rangle \rangle} \mathbb{E}_{J|I=i} [k_{IJ}] \right]^2 \\
&= \left[\sum_{i=1}^L \frac{n_i \langle w_{i\tilde{j}} \rangle}{M \langle \langle w_{i\tilde{j}} \rangle \rangle} (\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle) \right]^2 \\
&= (\langle \langle w_{i\tilde{j}} \rangle \rangle^{-1} \text{cov}_{\tilde{i}} [\langle w_{i\tilde{j}} \rangle, \langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle] + \text{ave}_{\tilde{i}} [\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle])^2
\end{aligned} \tag{S58}$$

Combining Eqs. (S54), (S55), and (S58), we obtain

$$\begin{aligned}
& \mathbb{E} \left[(\mathbb{E}_{J|I=i} [k_{IJ}] - \mathbb{E} [k_{IJ}])^2 \right] \\
&= \langle \langle w_{i\tilde{j}} \rangle \rangle^{-1} \text{cov}_{\tilde{i}} \left[\langle w_{i\tilde{j}} \rangle, (\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle)^2 \right] + \text{ave}_{\tilde{i}} \left[(\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle)^2 \right] \\
&\quad - (\langle \langle w_{i\tilde{j}} \rangle \rangle^{-1} \text{cov}_{\tilde{i}} [\langle w_{i\tilde{j}} \rangle, \langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle] + \text{ave}_{\tilde{i}} [\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle])^2 \\
&= \langle \langle w_{i\tilde{j}} \rangle \rangle^{-1} \text{cov}_{\tilde{i}} \left[\langle w_{i\tilde{j}} \rangle, (\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle)^2 \right] + \text{ave}_{\tilde{i}} \left[(\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle)^2 \right] \\
&\quad - (\langle \langle w_{i\tilde{j}} \rangle \rangle^{-1} \text{cov}_{\tilde{i}} [\langle w_{i\tilde{j}} \rangle, \langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle])^2 - (\text{ave}_{\tilde{i}} [\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle])^2 \\
&\quad - 2 \langle \langle w_{i\tilde{j}} \rangle \rangle^{-1} \text{cov}_{\tilde{i}} [\langle w_{i\tilde{j}} \rangle, \langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle] \text{ave}_{\tilde{i}} [\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle] \\
&= \langle \langle w_{i\tilde{j}} \rangle \rangle^{-1} \text{cov}_{\tilde{i}} \left[\langle w_{i\tilde{j}} \rangle, (\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle - \text{ave}_{\tilde{i}} [\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle])^2 \right] \\
&\quad - (\langle \langle w_{i\tilde{j}} \rangle \rangle^{-1} \text{cov}_{\tilde{i}} [\langle w_{i\tilde{j}} \rangle, \langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle])^2 \\
&\quad + \text{ave}_{\tilde{i}} \left[(\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle)^2 \right] - (\text{ave}_{\tilde{i}} [\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle])^2 \\
&= \langle \langle w_{i\tilde{j}} \rangle \rangle^{-1} \text{cov}_{\tilde{i}} \left[\langle w_{i\tilde{j}} \rangle, (\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle - \text{ave}_{\tilde{i}} [\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle])^2 \right] \\
&\quad - (\langle \langle w_{i\tilde{j}} \rangle \rangle^{-1} \text{cov}_{\tilde{i}} [\langle w_{i\tilde{j}} \rangle, \langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle])^2 \\
&\quad + \text{ave}_{\tilde{i}} \left[(\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle - \text{ave}_{\tilde{i}} [\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle])^2 \right]
\end{aligned} \tag{S59}$$

We consider each term in the last line of Eq. (S59) in terms of the order of s_a and s_w . We begin with the first term.

$$\begin{aligned}
& \langle \langle w_{i\tilde{j}} \rangle \rangle^{-1} \text{cov}_{\tilde{i}} \left[\langle w_{i\tilde{j}} \rangle, (\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle - \text{ave}_{\tilde{i}} [\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle])^2 \right] \\
& \langle \langle w_{i\tilde{j}} \rangle \rangle^{-1} \text{cov}_{\tilde{i}} \left[\langle w_{i\tilde{j}} \rangle, (\langle k_{i\tilde{j}} \rangle - \langle \langle k_{i\tilde{j}} \rangle \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle - \text{ave}_{\tilde{i}} [\Delta_s \langle k_{i\tilde{j}} \rangle])^2 \right] \\
&= \langle \langle w_{i\tilde{j}} \rangle \rangle^{-1} \text{cov}_{\tilde{i}} \left[\langle w_{i\tilde{j}} \rangle, (\langle k_{i\tilde{j}} \rangle - \langle \langle k_{i\tilde{j}} \rangle \rangle)^2 + (\Delta_s \langle k_{i\tilde{j}} \rangle - \text{ave}_{\tilde{i}} [\Delta_s \langle k_{i\tilde{j}} \rangle])^2 \right. \\
&\quad \left. - 2 (\langle k_{i\tilde{j}} \rangle - \langle \langle k_{i\tilde{j}} \rangle \rangle) (\Delta_s \langle k_{i\tilde{j}} \rangle - \text{ave}_{\tilde{i}} [\Delta_s \langle k_{i\tilde{j}} \rangle]) \right] \\
&= \langle \langle w_{i\tilde{j}} \rangle \rangle^{-1} \text{cov}_{\tilde{i}} \left[\langle w_{i\tilde{j}} \rangle, (\langle k_{i\tilde{j}} \rangle - \langle \langle k_{i\tilde{j}} \rangle \rangle)^2 \right] + O(s_w^2) \\
&\quad - 2 \langle \langle w_{i\tilde{j}} \rangle \rangle^{-1} \text{cov}_{\tilde{i}} \left[\langle w_{i\tilde{j}} \rangle, (\langle k_{i\tilde{j}} \rangle - \langle \langle k_{i\tilde{j}} \rangle \rangle) (\Delta_s \langle k_{i\tilde{j}} \rangle - \text{ave}_{\tilde{i}} [\Delta_s \langle k_{i\tilde{j}} \rangle]) \right],
\end{aligned} \tag{S60}$$

where we used Eqs. (S16) and (S17) in the last step. The last term of the last line of Eq. (S60) is zero because $\Delta_s \langle k_{i\tilde{j}} \rangle$ is independent of $\langle k_{i\tilde{j}} \rangle$ and $\langle w_{i\tilde{j}} \rangle$, a fact that stems from

the assumptions that $\partial s_w / \partial \langle k_{ij} \rangle = 0$ (see the main text under “Model”) and that v_{wi} is statistically uncorrelated with $\langle w_{i\tilde{j}} \rangle$ as i varies [see Eq. (S17)]. Thus, expanding $\langle w_{i\tilde{j}} \rangle$ as a Taylor series around $\langle k_{i\tilde{j}} \rangle = \langle\langle k_{i\tilde{j}} \rangle\rangle$, we can transform the last line of Eq. (S60) as follows:

$$\langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \text{cov}_{\tilde{i}} \left[\langle w_{i\tilde{j}} \rangle, (\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle)^2 \right] + O(s_w^2) = s_a c_a + O(s_w^2) + O(s_a^2), \quad (\text{S61})$$

where we introduced the following symbol:

$$c_a := \text{ave}_{\tilde{i}} \left[(\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle)^3 \right]. \quad (\text{S62})$$

Next, we consider the second term of the last line of Eq. (S59). Equation (S15) implies that

$$(\langle\langle w_{i\tilde{j}} \rangle\rangle^{-1} \text{cov}_{\tilde{i}} [\langle w_{i\tilde{j}} \rangle, \langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle])^2 = O(s_a^2). \quad (\text{S63})$$

Finally, we consider the third term of the last line of Eq. (S59) as follows:

$$\begin{aligned} & \text{ave}_{\tilde{i}} \left[(\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle - \text{ave}_{\tilde{i}} [\langle k_{i\tilde{j}} \rangle + \Delta_s \langle k_{i\tilde{j}} \rangle])^2 \right] \\ &= \text{ave}_{\tilde{i}} \left[(\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle + \Delta_s \langle k_{i\tilde{j}} \rangle - \text{ave}_{\tilde{i}} [\Delta_s \langle k_{i\tilde{j}} \rangle])^2 \right] \\ &= \text{ave}_{\tilde{i}} \left[(\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle)^2 + (\Delta_s \langle k_{i\tilde{j}} \rangle - \text{ave}_{\tilde{i}} [\Delta_s \langle k_{i\tilde{j}} \rangle])^2 \right. \\ &\quad \left. + 2 (\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle) (\Delta_s \langle k_{i\tilde{j}} \rangle - \text{ave}_{\tilde{i}} [\Delta_s \langle k_{i\tilde{j}} \rangle]) \right] \\ &= \text{ave}_{\tilde{i}} \left[(\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle)^2 \right] + \text{ave}_{\tilde{i}} \left[(\Delta_s \langle k_{i\tilde{j}} \rangle - \text{ave}_{\tilde{i}} [\Delta_s \langle k_{i\tilde{j}} \rangle])^2 \right] \\ &\quad + 2 \text{ave}_{\tilde{i}} \left[(\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle) (\Delta_s \langle k_{i\tilde{j}} \rangle - \text{ave}_{\tilde{i}} [\Delta_s \langle k_{i\tilde{j}} \rangle]) \right] \\ &= \text{ave}_{\tilde{i}} \left[(\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle)^2 \right] + \text{ave}_{\tilde{i}} \left[(\Delta_s \langle k_{i\tilde{j}} \rangle - \text{ave}_{\tilde{i}} [\Delta_s \langle k_{i\tilde{j}} \rangle])^2 \right] \\ &= \text{ave}_{\tilde{i}} \left[(\langle k_{i\tilde{j}} \rangle - \langle\langle k_{i\tilde{j}} \rangle\rangle)^2 \right] + O(s_w^2), \\ &= v_a + O(s_w^2), \end{aligned} \quad (\text{S64})$$

where we have assumed that $\langle k_{i\tilde{j}} \rangle$ and $\Delta_s \langle k_{i\tilde{j}} \rangle$ are statistically uncorrelated, an assumption that is essentially the same as the assumption made in Eq. (S17) that v_{wi} and $\langle w_{i\tilde{j}} \rangle$ are statistically uncorrelated.

Combining Eq. (S59), (S60), (S61), (S63), and (S64), we obtain

$$\mathbb{E} \left[(\mathbb{E}_{J|I=i} [k_{IJ}] - \mathbb{E} [k_{IJ}])^2 \right] = v_a + s_a c_a + O(s_a^2) + O(s_w^2). \quad (\text{S65})$$

Combining Eq. (S49), (S50), (S53), and (S65), we obtain

$$\mathbb{E} [v'_t] = (1 - M^{-1}) (v_a + s_a c_a + v_w - s_w c_w + m\sigma + O(s_a^2) + O(s_w^2)). \quad (\text{S66})$$

Substituting Eqs. (S48) and (S66) into $\mathbb{E} [v'_a] = \mathbb{E} [v'_t] - \mathbb{E} [v'_w]$, we obtain

$$\begin{aligned} \mathbb{E} [v'_a] &\approx (1 - M^{-1}) (v_a + s_a c_a + v_w - s_w c_w + m\sigma + O(s_a^2) + O(s_w^2)) \\ &\quad - (1 - \beta N^{-1}) (v_w + m\sigma - s_w c_w + O(s_w^2)) \\ &= (1 - M^{-1}) (v_a + s_a c_a + O(s_a^2) + O(s_w^2)) \\ &\quad + (\beta N^{-1} - M^{-1}) (v_w + m\sigma - s_w c_w + O(s_w^2)) \end{aligned} \quad (\text{S67})$$

4 Estimation of γ_a

Tsimring et al. [1] have investigated the time evolution of the probability density $p(r, t)$ of fitness r subject to mutation and selection. In this section, we show that the results of Tsimring et al. [1] imply $C \approx -0.25V^{3/2}$, where V and C are the variance and the third central moment of $p(r, t)$, respectively. This implication is consistent with our postulate $c_a = -\gamma_a v_a^{3/2}$ made in Eq. (9) of the main text, where γ_a was measured to be about 0.25 through simulations.

Tsimring et al. [1] have considered the following equation, which describes the time evolution of $p(r, t)$:

$$\frac{\partial}{\partial t} p(r, t) = \theta(p - p_c) (r - \langle r \rangle) p(r, t) + D \frac{\partial^2}{\partial r^2} p(r, t), \quad (\text{S68})$$

where $\theta(x)$ is the Heaviside step function, and $\langle r \rangle$ is the average fitness defined as

$$\langle f(r) \rangle = \int_{-\infty}^{\infty} f(r) p(r, t) dr,$$

and D is a diffusion constant. The first term on the RHS of Eq. (S68) describes the effect of selection; the second term, that of mutation. The Heaviside step function accounts for the fact that the probability density $p(r, t)$ must exceed a small threshold density p_c to grow because the size of a population is not infinite in reality. Tsimring et al. have shown that Eq. (S68) allows a travelling-wave solution, in which the peak of the density travels toward higher values of r , while maintaining a pulse-like shape, at a steady-state speed (denoted by v)

$$v = cD^{2/3}, \quad (\text{S69})$$

where the value of c depends weakly on p_c and is around 4 in a wide range of p_c [1].

Multiplying both sides of Eq. (S68) with r or $(r - \langle r \rangle)^2$ and integrating over the whole range, we get

$$\frac{d}{dt} \langle r \rangle = V - \epsilon_1, \quad (\text{S70})$$

$$\frac{d}{dt} V = C - \epsilon_2 + 2D, \quad (\text{S71})$$

where V , C , ϵ_1 , and ϵ_2 are defined as follows:

$$V = \langle (r - \langle r \rangle)^2 \rangle, \quad (\text{S72})$$

$$C = \langle (r - \langle r \rangle)^3 \rangle, \quad (\text{S73})$$

$$\epsilon_1 = \langle \theta(p_c - p) (r - \langle r \rangle)^2 \rangle, \quad (\text{S74})$$

$$\epsilon_2 = \langle \theta(p_c - p) (r - \langle r \rangle)^3 \rangle. \quad (\text{S75})$$

In obtaining Eqs. (S70) and (S71), we have assumed that the surface terms go to zero as $r \rightarrow \pm\infty$; i.e., $\lim_{r \rightarrow \pm\infty} P(r, t) = 0$, $\lim_{r \rightarrow \pm\infty} rP(r, t) = 0$, $\lim_{r \rightarrow \pm\infty} r \frac{\partial p}{\partial r} = 0$, and $\lim_{r \rightarrow \pm\infty} r^2 \frac{\partial p}{\partial r} = 0$.

For a travelling-wave solution of Eq. (S68) with a constant speed v and shape, Eq. (S70) implies

$$v = V - \epsilon_1. \quad (\text{S76})$$

Table S1: **Correspondence between Kimura's notation [2] and ours.**

Kimura's	ours	description
c	s_a	among-collective selection coefficient
v	m	mutation rate per generation from non-altruistic to altruistic allele
v'	m	reverse mutation rate; we assumed $v = v'$
s'	s_w	within-collective selection coefficient
m	0	among-collective migration rate
$2N$	$\beta^{-1}N$	number of alleles per collective; Kimura considers diploid
∞	M	total number of alleles

From Eqs. (S69) and (S76), we get

$$D = \left(\frac{V - \epsilon_1}{c} \right)^{3/2}. \quad (\text{S77})$$

Since v and ϵ_1 are constant, Eq. (S76) implies $dV/dt = 0$. Thus, Eq. (S71) implies

$$C = -2D + \epsilon_2. \quad (\text{S78})$$

Equations (S77) and (S78) imply

$$C = -2 \left(\frac{V - \epsilon_1}{c} \right)^{3/2} + \epsilon_2 \approx -2c^{-3/2}V^{3/2},$$

where we have assumed $\epsilon_1 \ll V$ and $\epsilon_2 \ll V$ to obtain the last term. Since c is about 4 according to Tsimring et al. [1], we get

$$C \approx -0.25V^{3/2}.$$

5 Converting Kimura's notation into ours

Kimura [2] has investigated a binary-trait model of multilevel selection and shown that within-collective selection exactly balances out among-collective selection if

$$\frac{c}{v + v' + m} - 4Ns' = 0, \quad (\text{S79})$$

where the symbols are as described in Table S1 (see also the next paragraph). Equation (S79) includes Eq. (17) of the main text as a special case. Equation (S79) appears as Eq. (27) of Ref. [2] or Eq. (4.8) of Ref. [3] as and is derived therein under the assumption that the steady-state frequency of the altruistic allele is identical to that in the absence of selection, an approximation that is expected to be valid in the limit of weak selection.

To convert Kimura's notation into ours, we assumed that the rate of mutation from a non-altruistic to an altruistic allele is identical to the rate of mutation from the altruistic to the non-altruistic allele, so that the direction of mutation is unbiased as in our quantitative-trait model. Moreover, we assumed that the migration rate among collectives is zero since

our model does not consider migration. Finally, we took account of the fact that Kimura's model considers diploid as follows. In Kimura's model, each collective consists of N diploid individuals, i.e., $2N$ alleles. The number of alleles per collective can be considered as the average number of replicators per collective in our model (i.e., $\beta^{-1}N$) because Kimura's model assumes no dominance.

6 Derivation of Kimura's result through our method

In this section, we derive Eq. (17) of the main text, which gives parameter-region boundaries of the binary-trait model, using the method developed in the main text. The most important difference between the binary-trait and quantitative-trait models resides in the definition of ϵ_{IJ} . Thus, we consider only terms involving ϵ or $m\sigma$ as described below.

First, we show that the change in the definition of ϵ_{IJ} does not affect the condition for the parameter-region boundary given by Eq. (3). In the binary-trait model, $\epsilon_{IJ} = 0$ with probability $1 - m$ and $\epsilon_{IJ} = 1 - 2k_{IJ}$ with probability m (I and J are random variables taking the indices of a sampled replicator, as defined in Section 1). Thus,

$$\begin{aligned}\mathbb{E}[\epsilon_{IJ}] &= \sum_{i=1}^L \sum_{j=1}^{n_i} P(I=i, J=j) \int dP(\epsilon_{IJ}) \epsilon_{IJ} \\ &= \sum_{i=1}^L \sum_{j=1}^{n_i} \frac{w_{ij}}{M \langle\langle w_{ij} \rangle\rangle} m(1 - 2k_{ij}) \\ &= m(1 - 2\langle\langle k_{ij} \rangle\rangle) + O(s_w^2 + s_a^2 + ms_w + ms_a).\end{aligned}\tag{S80}$$

Therefore, Eq. (4) needs to be modified as follows:

$$\mathbb{E}[\Delta \langle\langle k_{ij} \rangle\rangle] = s_a v_a - s_w v_w + m(1 - 2\langle\langle k_{ij} \rangle\rangle) + O(s_w^2 + s_a^2 + ms_w + ms_a).\tag{S81}$$

This equation, however, becomes almost identical to Eq. (4) if the parameters are on the parameter-region boundary, on which $\langle\langle k_{ij} \rangle\rangle = 1/2$. Thus, the condition for the parameter-region boundary when s_a , s_w , and m are sufficiently small is the same as in the quantitative-trait model.

Next, we consider Eq. (5) and show that the change in the definition of ϵ_{IJ} makes a significant difference, which explains the difference between the binary- and quantitative-trait models. In Eq. (5), $m\sigma$ represents the difference between v_t and the variance of $k_{IJ} + \epsilon_{IJ}$. In the quantitative-trait model, this difference is simply the variance of ϵ_{IJ} because ϵ_{IJ} and k_{IJ} are independent of each other. In the binary-trait model, however, ϵ_{IJ} and k_{IJ} are not independent, and this fact affects Eq. (5), as follows. Under the assumption that $s_a = s_w = 0$, the variance of $k_{IJ} + \epsilon_{IJ}$ in the binary-trait model is

$$\begin{aligned}\mathbb{E}[(k_{IJ} + \epsilon_{IJ} - \mathbb{E}[k_{IJ} + \epsilon_{IJ}])^2] &= (1 - \mathbb{E}[k_{IJ} + \epsilon_{IJ}])^2 P(k_{IJ} + \epsilon_{IJ} = 1) \\ &\quad + (0 - \mathbb{E}[k_{IJ} + \epsilon_{IJ}])^2 P(k_{IJ} + \epsilon_{IJ} = 0),\end{aligned}\tag{S82}$$

where

$$\begin{aligned}P(k_{IJ} + \epsilon_{IJ} = 1) &= \langle\langle k_{ij} \rangle\rangle(1 - m) + (1 - \langle\langle k_{ij} \rangle\rangle)m \\ P(k_{IJ} + \epsilon_{IJ} = 0) &= (1 - \langle\langle k_{ij} \rangle\rangle)(1 - m) + \langle\langle k_{ij} \rangle\rangle m \\ \mathbb{E}[k_{IJ} + \epsilon_{IJ}] &= P(k_{IJ} + \epsilon_{IJ} = 1).\end{aligned}\tag{S83}$$

Thus,

$$\mathbb{E}[(k_{IJ} + \epsilon_{IJ} - \mathbb{E}[k_{IJ} + \epsilon_{IJ}])^2] = v_t + m(1-m)(1 - 2\langle\langle k_{ij} \rangle\rangle)^2, \quad (\text{S84})$$

where we used the fact that $v_t = \langle\langle k_{ij} \rangle\rangle(1 - \langle\langle k_{ij} \rangle\rangle)$. Therefore, the expected sample variance of the next generation is

$$\mathbb{E}[v'_t] = (1 - M^{-1})[v_t + m(1-m)(1 - 2\langle\langle k_{ij} \rangle\rangle)^2]. \quad (\text{S85})$$

Likewise, under the assumption that all collectives always consist of $\beta^{-1}N$ replicators, the expected sample variance within a collective of the next generation is

$$\mathbb{E}[v'_{wi}] = (1 - \beta N^{-1})[v_{wi} + m(1-m)(1 - 2\langle k_{ij} \rangle)^2], \quad (\text{S86})$$

where the index of collectives i needs to be kept because $\langle k_{ij} \rangle$ depends on i . Averaging $\mathbb{E}[v'_{wi}]$ over i , we obtain

$$\begin{aligned} \mathbb{E}[v'_w] &\approx \text{ave}_i[\mathbb{E}[v'_{wi}]] \\ &= (1 - \beta N^{-1})\{v_w + m(1-m)[(1 - 2\langle\langle k_{ij} \rangle\rangle)^2 + 4v_a]\}, \end{aligned} \quad (\text{S87})$$

where we used the fact that $v_a = \text{ave}_i[\langle k_{ij} \rangle^2] - \langle\langle k_{ij} \rangle\rangle^2$. Since $\mathbb{E}[v'_a] = \mathbb{E}[v'_t] - \mathbb{E}[v'_w]$, we obtain

$$\begin{aligned} \mathbb{E}[v'_a] &= (1 - M^{-1})v_a + (\beta N^{-1} - M^{-1})\left[v_w + m(1-m)(1 - 2\langle\langle k_{ij} \rangle\rangle)^2\right] \\ &\quad - 4(1 - \beta N^{-1})m(1-m)v_a. \end{aligned} \quad (\text{S88})$$

If the systems is on a parameter-region boundary, $\langle\langle k_{ij} \rangle\rangle = 1/2$. Thus, setting $\langle\langle k_{ij} \rangle\rangle = 1/2$, we obtain

$$\mathbb{E}[v'_w] = (1 - \beta N^{-1})[v_w + 4m(1-m)v_a] \quad (\text{S89})$$

$$\mathbb{E}[v'_a] = (1 - M^{-1})v_a + (\beta N^{-1} - M^{-1})v_w - 4(1 - \beta N^{-1})m(1-m)v_a. \quad (\text{S90})$$

To apply the condition for the parameter-region boundary $v_w/v_a \approx s_a/s_w$, we need to calculate the ratio v_w/v_a . To this end, dividing Eq. (S89) by Eq. (S90) on each side, we obtain

$$\frac{\mathbb{E}[v'_w]}{\mathbb{E}[v'_a]} = \frac{(1 - \beta N^{-1})\left[\frac{v_w}{v_a} + 4m(1-m)\right]}{(1 - M^{-1}) + (\beta N^{-1} - M^{-1})\frac{v_w}{v_a} - 4(1 - \beta N^{-1})m(1-m)}. \quad (\text{S91})$$

Assuming a steady state (i.e., $\mathbb{E}[v'_w]/\mathbb{E}[v'_a] = v_w/v_a$), we obtain

$$\frac{v_w}{v_a} = \frac{4m(1-m)(1 - \beta N^{-1})}{\beta N^{-1} - M^{-1}}. \quad (\text{S92})$$

Using the condition $v_w/v_a \approx s_a/s_w$, we obtain

$$\frac{4m(1-m)(1 - \beta N^{-1})}{\beta N^{-1} - M^{-1}} \approx \frac{s_a}{s_w}. \quad (\text{S93})$$

If $\beta N^{-1} \ll 1$, $m \ll 1$, and $M \rightarrow \infty$ (Kimura's model assumes that $M = \infty$), we obtain

$$\frac{s_a}{s_w} \approx 4m\beta^{-1}N, \quad (\text{S94})$$

which is the same as Eq. (17).

7 Supplementary Figures

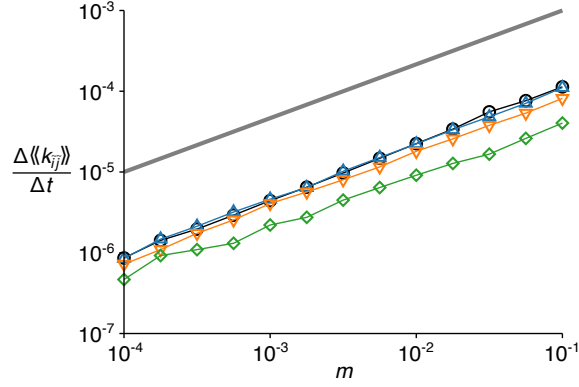


Figure S1: Rate of logarithmic fitness increase as function of mutation rate measured through simulations with no within-collective selection ($s_w = 0$, $s_a = 10^{-3}$, $M = 5 \times 10^5$, and $\sigma = 10^{-4}$). Fitness is defined as $w_{ij} = e^{s_a \langle k_{ij} \rangle}$. Symbols have following meaning: $N = 10^2$ (black circles); $N = 10^3$ (blue triangle up); $N = 10^4$ (orange triangle down); $N = 10^5$ (green diamond). Line is $\Delta \langle \langle k_{ij} \rangle \rangle / \Delta t \propto m^{2/3}$, as predicted by Eqs. (4) and (14) in main text. This figure confirms that $\Delta \langle \langle k_{ij} \rangle \rangle / \Delta t \propto m^{2/3}$ in agreement with Ref. [1]. Note also that $\Delta \langle \langle k_{ij} \rangle \rangle$ is roughly independent of N if $N \ll M$, which is consistent with prediction of Eq. (14) in main text.

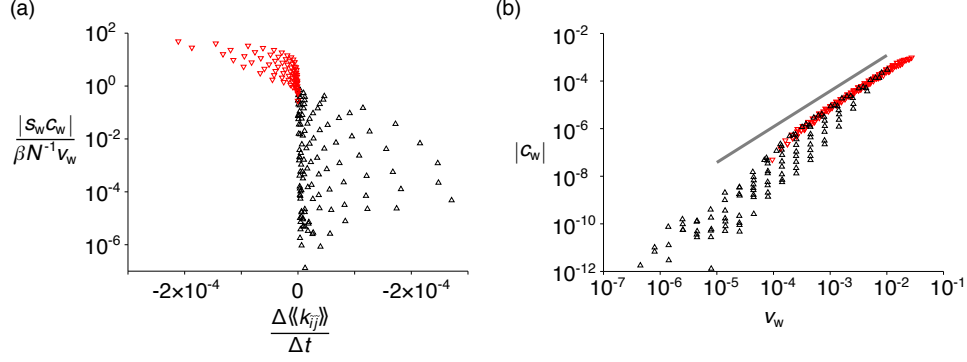


Figure S2: Average third central moment of k_{ij} within collective (c_w) measured through simulations ($M = 5 \times 10^5$, $\sigma = 10^{-4}$, and $s_w = s_a = 10^{-2}$). (a) Ratio between effect of selection and that of random genetic drift on Δv_w as function of $\Delta \langle k_{ij} \rangle / \Delta t$: $\Delta \langle k_{ij} \rangle / \Delta t > 0$ (black triangle up); $\Delta \langle k_{ij} \rangle / \Delta t < 0$ (red triangle down). (b) $|c_w|$ as function of v_w . Triangles are simulation results: $\Delta \langle k_{ij} \rangle / \Delta t > 3 \times 10^{-7}$ (black triangle up); $\Delta \langle k_{ij} \rangle / \Delta t < 3 \times 10^{-7}$ (red triangle down). Line is $|c_w| \propto v_w^{3/2}$, as postulated in Eq.(9). Least squares fitting of $|c_w| = \gamma_w v_w^{3/2}$ to data for $\Delta \langle k_{ij} \rangle / \Delta t < 3 \times 10^{-7}$ yielded $\gamma_w \approx 0.25$.

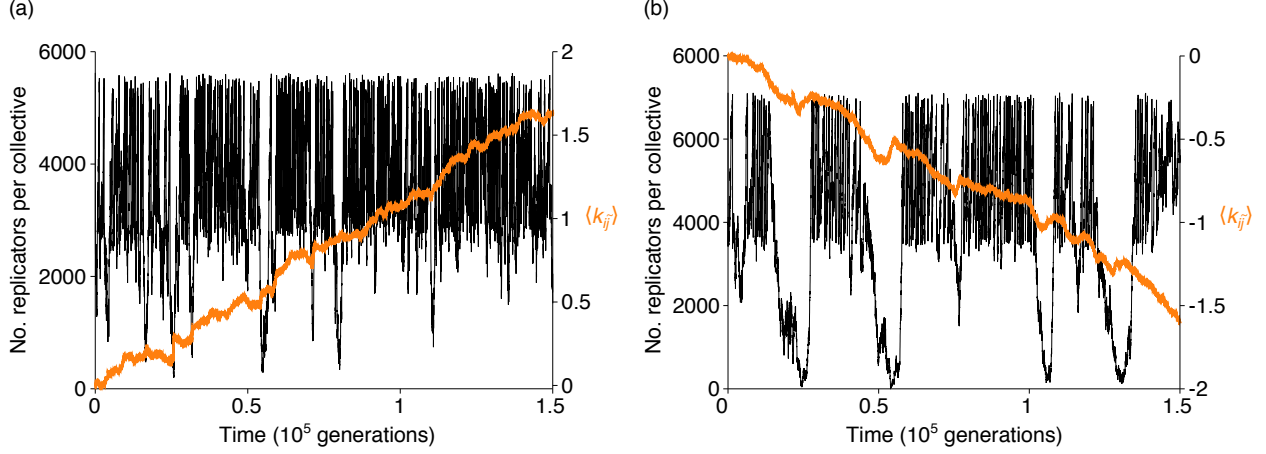


Figure S3: Dynamics of common ancestors of collectives ($m = 0.01$, $M = 5 \times 10^5$, $\sigma = 10^{-4}$, and $s_a = s_w = 0.01$). Plotted are number of replicators per collective (black; left coordinate) and $\langle k_{ij} \rangle$ (orange; right coordinate). (a) $N = 5623$. In this case, $\Delta \langle k_{ij} \rangle > 0$, and evolutionarily stable disequilibrium is not clearly observed. (b) $N = 17783$. In this case, $\Delta \langle k_{ij} \rangle < 0$, and evolutionarily stable disequilibrium is clearly observed.

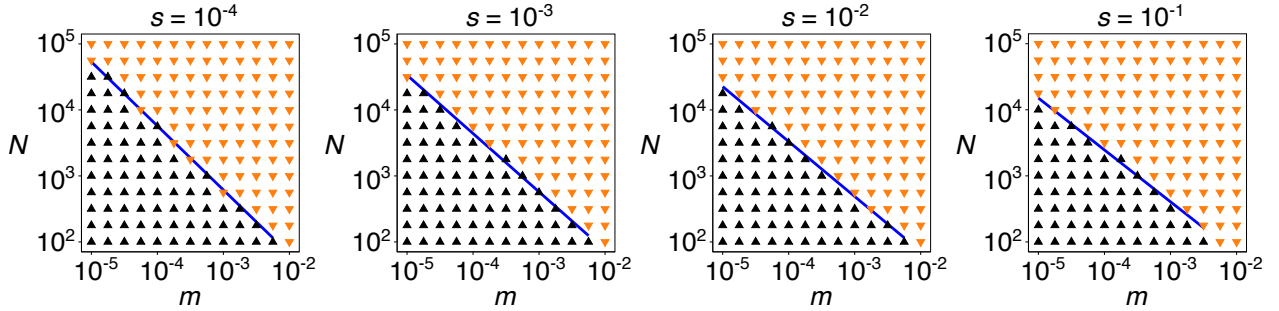


Figure S4: Parameter-sweep diagrams of binary-trait model ($s_w = s_a = s$ and $M = 5 \times 10^5$). Symbols have following meaning: $\langle k_{ij} \rangle > 1/2$ (triangle up); $\langle k_{ij} \rangle < 1/2$ (triangle down). Lines are estimated parameter-region boundaries. Parameter-region boundaries were estimated as follows. Zeros of $\langle k_{ij} \rangle - 1/2$ were estimated with linear interpolation with respect to N from two simulation points around parameter-region boundary for various m values between 10^{-5} and 10^{-1} . Estimated zeros were used to obtain parameter-region boundary through least squares regression of $N \propto m^{-\alpha}$.

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